

IDENTIFICATION OF SIMULTANEOUS EQUATIONS MODEL

**ECONOMETRICS - II
COMPUTER PROJECT**

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OBJECTIVE

The objective of the project is to demonstrate the identification of simultaneous equations model for sales and advertising relationships between filter and non-filter cigarette brands. The model tests for the success of the participating tobacco companies in targeting the youth in the United States.

Motivation

Advertising is an important method of competition in industries that are highly concentrated, such as the cigarette industry. Firms in industries of this type tend not to compete by price, but try to increase sales with advertising and other marketing techniques. Since a number of prior studies have found little relationship between advertising and tobacco consumption, it is important to examine this literature before proceeding with a new empirical study. Economic theory provides some insights into how econometric studies of tobacco advertising should be conducted. An important economic aspect of advertising is diminishing marginal product. Diminishing marginal product suggests that, after some point, additions to advertising will result in ever smaller additions to consumption. This concept is the basis of the advertising response function. Advertising response functions have been used for some time in brand level research to illustrate the effect of advertising at various levels on consumption.

Many of the recent advertising exposure and recall studies have controlled for other risk factors for smoking initiation. This means that the alleged causal factor of advertising should be isolated from other risk factors so that a genuine association can be determined. For instance, it might be true that 16-year-olds who can remember tobacco advertisements are more likely to smoke as 20-year-olds than 16-year-olds who do not remember tobacco advertisements, but it does not follow from this that this association is due to the fact that they remembered tobacco advertisements. It might well be that those 16-year-olds might also have shared some other characteristic such as being more inclined to risk-taking, or performing poorly at school, and it may be that these factors - which have been excluded from most studies - rather than tobacco advertisements, are the cause(s) of subsequent smoking. The point is that without controlling for these other factors one can never know.

Many of the recall studies are rendered problematic by the fact that they assume that the direction of influence between advertising and smoking uptake and consumption is one-way. That is, they assume that the reputedly causal direction is from exposure to advertising. Yet there is very little empirical support offered for this. Indeed, important studies have contradicted it. Adolescents who are interested in and receptive to smoking can be assumed to have a smoking preference that would lead to their interest in and recall of tobacco advertising. Their preference for smoking could have been formed prior to and independently of any tobacco advertisement. This would mean that discovering an association

between remembering tobacco advertisements and subsequently smoking does not constitute compelling evidence that the former leads to the latter.

Moreover, tobacco advertising exposure and recall studies often rely on self-reported data, which poses a significant problem in terms of validity. For instance, self-reported data may be inaccurate because of certain situational factors associated with the social desirability of the behaviour, such as smoking, being studied. The research literature notes that self-reports are often plagued by social desirability bias, that is, subjects often report what they believe the interviewer wishes to hear or what is the socially preferred option, rather than what is actually the case. Some studies have shown that the likelihood of untruthful responses rises with the degree of threat posed by the question. But even setting aside the question of conscious deception, there is the possibility of significant inaccuracy due to faulty recall.

Many tobacco advertising exposure and recall studies suffer from a variety of methodological issues which undermine their internal validity. These include misspecified variances, omitted interactions and paths, endogeneity, sample attrition and selection bias. As noted recently by Nelson, most longitudinal studies of advertising exposure and recall fail to deal with these issues. (*J. Nelson What is learned from longitudinal studies of advertising and youth drinking and smoking? A critical assessment International Journal of Environmental Research and Public Health*20107:870-926.

Identification: A historical perspective

Intriligator (1978) defined identification as 'the problem of relating the structural parameters of a simultaneous equation model to the reduced-form parameters that "summarize all relevant information available from the sample data"'. However, Bartel (1985) was the first to attempt a heuristic definition of identification, where he claimed that identification was a problem that related directly to probabilistic distributions. Identification has also been done in a Bayesian framework and for triangular simultaneous equation systems for control variables (Imben and Newey, 2009). An interesting practical application of the concept of identification has been in the case of standard auction models. Athey and Haile (2002) present identification results for models of first-price, second-price, ascending and descending auctions. Rigobon (2003) suggested an alternative method of identification by using the heteroscedasticity present in the model. This method was used to measure the contemporaneous propagation between the returns on several Latin American sovereign bonds. Berry formulated an alternative algorithm to solve for identification, which is equivalent to the rank condition and simpler to evaluate.

Identification in simultaneous equations model:

A simultaneous equations system is defined as a system with two or more equations, where a variable explained in one equation appears as an explanatory variable in another. Thus, the endogenous variables in the system are simultaneously determined.

The Structural Form:

The structural form of the simultaneous equations system can be written as follows:

$$\beta' Y_t + \Gamma' X_t = \varepsilon_t$$

Where:

- β' - matrix of coefficients of endogenous variables of order $(n \times n)$
- Y_t - vector of endogenous variables of order $(n \times 1)$
- Γ' - matrix of coefficients of exogenous variables of order $(n \times m)$
- X_t - vector of exogenous variables of order $(m \times 1)$
- ε_t - vector of error terms of order $(n \times 1)$

The equations above are known as Structural or Behavioral Equations because they portray the structure of an economy or the behavior of an economic agent. The β' and Γ' are known as **structural parameters** or **coefficients**.

The Reduced Form:

The reduced form for simultaneous equations system can be obtained as follows:

We know,

$$\beta' Y_t + \Gamma' X_t = \varepsilon_t$$

This can be rewritten as:

$$\beta' Y_t = -\Gamma' X_t + \varepsilon_t$$

Since β' is a non-singular matrix [i.e. $\det(\beta')$ is not equal to zero], it is invertible, thus:

$$Y_t = -(\beta')^{-1} \Gamma' X_t + (\beta')^{-1} \varepsilon_t$$

$$Y_t = \pi' X_t + u_t$$

Where:

$$\diamond \pi' = -(\beta')^{-1} \Gamma'$$

$$\diamond u_t = (\beta')^{-1} \varepsilon_t$$

A reduced form equation expresses an endogenous variable solely in terms of the predetermined variables and the stochastic disturbances. The reduced form coefficients are also known as **Impact** or **Short Run Multipliers**.

The Question of Identification:

By identification we mean whether numerical estimates of the parameters of a structural can be obtained from the estimated reduced form coefficients. If this can be done, we say that the particular equation is *identified* else the equation under consideration is *over-identified* or *under-identified*. Thus, if we evaluate π and the relation between π , β and Γ , then identification poses the question of whether there exists a unique solution for β_1 and Γ_1 (the parameters of the first equation)?

Identification is checked for the structural parameters, equation by equation. There are two alternative approaches to study this:

- By a study of the linear combinations of the equations of the system, that is, from the structural form.
- By a study of the relation between π , β and Γ .

A structural equation is said to be identified if and only if all its parameters are identified. The parameters of a structural equation are said to be identified if and only if they can be obtained in a unique way from the reduced form. **A simultaneous equations model (system) is said to be identified if and only if all the structural equations in the system are identified.**

If there are no a priori restrictions on β_1 and Γ_1 , then the equation has no solution and therefore cannot be identified. A priori restrictions on the model enable identification of the structural equations. There are two kinds of restrictions:

- **Normalization restrictions:** Such a restriction requires one of the coefficients to be reduced to 1 or -1.
- **Exclusion restrictions:** Such a restriction requires one of the coefficients to be reduced to zero.

Rank and Order Conditions for Identification:

Rank Condition:

To derive a unique solution for β^* from the reduced form model:

$$r(\pi_{\Delta\Delta^*}) = r[\pi_{\Delta\Delta^*} \quad \pi_{\Delta\Delta 1}] = n^*$$

This condition is both a necessary as well as a sufficient condition.

Order Condition:

For the rank condition to hold the following order condition is a necessary condition:

$$m_{\Delta\Delta} \geq n^*$$

Where:

- $m_{\Delta\Delta}$ represents the number of excluded exogenous variables from the equation.
- n^* represents the number of included endogenous explanatory variables.

This implies that there should be sufficient enough exogenous variables to control for the simultaneously changing endogenous variables. This implies that:

$$m_{\Delta\Delta} + m_{\Delta} \geq n^* + m_{\Delta} \Rightarrow m \geq n^* + m_{\Delta}$$

- If $m_{\Delta\Delta} < n^* \Rightarrow$ Equation is not identified
- If $m_{\Delta\Delta} = n^* \Rightarrow$ Equation is exactly identified (assuming that the rank condition is satisfied)

If $m_{\Delta\Delta} > n^* \Rightarrow$ Equation is overidentified (assuming that the rank condition is satisfied)

Reduced form Identification:

To determine the identification of the first equation in the simultaneous equation system:

Step1: Let Y_1 be the explained variable in the first equation, normalize the coefficient attached to this variable equal to -1.

Step 2: Let there be n^* endogenous explanatory variables in the first equation ($Y_2, Y_3, Y_4, \dots, Y_{n^*+1}$).

$$\text{Partition } Y = [Y_1 \ Y^* \ Y^{**}]$$

Where:

- Y_1 is the $(T \times 1)$ matrix of the explained variable.
- Y^* is the $(T \times n^*)$ matrix of endogenous explanatory variables that appear in the first equation.
- Y^{**} is the $(T \times n^{**})$ matrix of endogenous variables that do not appear in the first equation.

Total number of endogenous variables: $n = 1 + n^* + n^{**}$

Step3: Let there be m_Δ exogenous explanatory variables in the first equation:

$$\text{Partition } X = [X_\Delta \ X_{\Delta\Delta}]$$

Where:

- X_Δ is the $(T \times m_\Delta)$ matrix of exogenous explanatory variables that appear in the first equation.
- $X_{\Delta\Delta}$ is the $(T \times m_{\Delta\Delta})$ matrix of exogenous variables that do not appear in the first equation.

Total number of exogenous variables: $m = m_\Delta + m_{\Delta\Delta}$

Step 4: Now partition β_1 and Γ_1 accordingly:

$$\beta_1 = \begin{pmatrix} -1 \\ \beta^* \\ 0 \end{pmatrix}$$

$$\Gamma_1 = \begin{pmatrix} \Gamma_\Delta \\ 0 \end{pmatrix}$$

Step 5: Now the first equation, $Y\beta_1 + X\Gamma_1 = \varepsilon_1$ can be written as:

$$[Y_1 \ Y^* \ Y^{**}] \begin{pmatrix} -1 \\ \beta^* \\ 0 \end{pmatrix} + [X_\Delta \ X_{\Delta\Delta}] \begin{pmatrix} \Gamma_\Delta \\ 0 \end{pmatrix} = \varepsilon_1$$

$$\Leftrightarrow -Y_1 + Y^*\beta^* + X_\Delta\Gamma_\Delta = \varepsilon_1$$

Step 6: Consider the reduced form: $\pi\beta = -\Gamma_1$

$$\begin{matrix} \pi_{\Delta 1} & \pi_{\Delta^*} & \pi_{\Delta^{**}} \\ \pi_{\Delta\Delta 1} & \pi_{\Delta\Delta^*} & \pi_{\Delta\Delta^{**}} \end{matrix} \begin{pmatrix} -1 \\ \beta^* \\ 0 \end{pmatrix} = \begin{pmatrix} \Gamma_{\Delta} \\ 0 \end{pmatrix}$$

$$-\pi_{\Delta 1} + \pi_{\Delta^*} \beta^* = -\Gamma_{\Delta} \quad (1)$$

$$-\pi_{\Delta\Delta 1} + \pi_{\Delta\Delta^*} \beta^* = 0 \quad (2)$$

If there is a unique solution of β^* from equation (2), then we can get a unique solution of Γ_{Δ} from equation (1).

Therefore, the problem of identification reduces to determining the conditions which give a unique solution of β^* from equation (2).

$$\pi_{\Delta\Delta^*} \beta = \pi_{\Delta\Delta 1}$$

Structural form Identification:

Consider a Simultaneous Equations System with 4 endogenous variables and 3 exogenous variables:

$$Y_{1t} - \beta_{10} - \beta_{12}Y_{2t} - \beta_{13}Y_{3t} - \gamma_{11}X_{1t} = u_{1t}$$

$$Y_{2t} - \beta_{20} - \beta_{23}Y_{3t} - \gamma_{21}X_{1t} - \gamma_{22}X_{2t} = u_{2t}$$

$$Y_{3t} - \beta_{30} - \beta_{31}Y_{1t} - \gamma_{31}X_{1t} - \gamma_{32}X_{2t} = u_{3t}$$

$$Y_{4t} - \beta_{40} - \beta_{41}Y_{1t} - \beta_{42}Y_{2t} - \gamma_{43}X_{3t} = u_{4t}$$

Using the **Order Condition**, we verify that each equation is exactly identified:

No. of predetermined variables excluded, $(K - k)$	No. of endogenous variables included less one, $(m - 1)$	Identified?
2	2	Exactly
1	1	Exactly
1	1	Exactly
2	2	Exactly

Let us now verify the **Rank Condition** for Structural Equations:

“In a model containing M equations in M endogenous variables, an equation is identified if and only if *at least* one non zero determinant of order (M-1)x(M-1) can be constructed from the coefficients of the variables (both endogenous and predetermined) excluded from that particular equation but included in the other equations of the model.”

Coefficients of the variables

1	Y_1	Y_2	Y_3	Y_4	X_1	X_2	X_3
$-\beta_{10}$	1	$-\beta_{12}$	$-\beta_{13}$	0	$-\gamma_{11}$	0	0
$-\beta_{20}$	0	1	$-\beta_{23}$	0	$-\gamma_{21}$	$-\gamma_{22}$	0
$-\beta_{30}$	$-\beta_{31}$	0	1	0	$-\gamma_{31}$	$-\gamma_{32}$	0
$-\beta_{40}$	$-\beta_{41}$	$-\beta_{42}$	0	1	0	0	$-\gamma_{43}$

Consider the first equation which excludes the variables Y_4, X_2, X_3 . For this equation to be identified we must obtain at least one non-zero determinant of order 3x3 from the coefficients of the variables excluded from this equation but included in other equations. The rank condition of the other equations will also follow in the same manner. In case, such a matrix is not found or the determinant of the matrix is equal to zero, the particular equation is unidentified. Consequently, the system of equations will also be unidentified.

The Model:

The model studies the expenditure on cigarette advertisements in national newspapers, magazines and promotional billboards in USA over a period of 1952- 1965. The model focuses on those modes of advertisement which are closely targeted by the population above the age of 20 years.

The model consists of two demand equations for two competing groups of cigarette brands and two equations that describe the advertising relations of these groups of brands. The sales of the major filter cigarette brands have been aggregated to give one demand equation for this group. Similarly, there is one demand equation for the major non-filter brands.

Since the prices of filter brands are identical as are the prices of non-filter brands, the aggregation of brand sales in each class is justified theoretically. The Leontieff-Hick’s Theorem establishes that if the prices of a group of goods change in equal proportion, that group can then be treated as a single commodity. Since, this theorem justifies aggregation in this study; the concepts of complementarity,

substitutability, price elasticity, income elasticity apply to the grouped commodities just as the corresponding concepts and measures apply to single goods.

Variables used in the model:

- $\text{Log}(\text{Sales}_t)$: Logarithm of sales for filter cigarettes divided by population over age 20.
- $\text{Log}(\text{Adf}_t)$: Logarithm of advertising dollar sales for filter cigarettes divided by population over age 20.
- $\text{Log}(\text{Adnf}_t)$: Logarithm of advertising dollar sales for non-filter cigarettes divided by population over age 20.
- $\text{Log}(\text{PDI}_t)$: Logarithm of disposable personal income divided by population over age 20 divided by consumer price index.
- $\text{Log}(\text{Price}_t)$: Logarithm of price per package of non-filter divided by consumer price index.
- $\text{Log}(\text{Salesnf}_t)$: Logarithm of sales for non-filter cigarettes divided by population over age 20.

Demand for Major Brands of Filter and Non-Filter Cigarettes:

Demand Equation for Filter Brands:

For every year t , demand for filter brands is considered in isolation from the rest of the system,

$$\text{Log}(\text{Sales}_t) = \beta_1 \text{Log}(\text{Adf}_t) + \beta_2 \text{Log}(\text{Adnf}_t) + \gamma_1 \text{Log}(\text{PDI}_t) + \gamma_2 \text{Log}(\text{Price}_t) + \varepsilon_t$$

We therefore postulate that the per capita sales of filter cigarettes is a non-linear function of the ratio of per capita advertising for the two competitive types of cigarettes and the two exogenous variables. Although it might have been desirable to include prices of the filter and non-filter cigarettes as variables, the non-filter price is available as a component of CPI but the filter price is not.

Demand Equation for Non-Filter Brands:

For every year t , demand for Non-filter brands is considered in isolation from the rest of the system,

$$\text{Log}(\text{Salesnf}_t) = \beta_3 \text{Log}(\text{Adf}_t) + \beta_4 \text{Log}(\text{Adnf}_t) + \gamma_4 \text{Log}(\text{PDI}_t) + \gamma_5 \text{Log}(\text{Price}_t) + \varepsilon_{2t}$$

Advertising Relationships for Major Brands of Filter and Non-Filter Cigarettes:

Equation describing Advertising Behaviour of Filter Brands:

$$\text{Log}(\text{Salesf}_t) = \beta_5 \text{Log}(\text{Salesnf}_t) + \beta_6 \text{Log}(\text{Adf}_t) + \epsilon_{3t}$$

Equation describing Advertising Behaviour of Non-Filter Brands:

$$\text{Log}(\text{Salesnf}_t) = \beta_7 \text{Log}(\text{Salesf}_t) + \beta_6 \text{Log}(\text{Adnf}_t) + \epsilon_{4t}$$

Model of Sales and Advertising of Filter and Non-Filter Cigarettes:

The model's parts produce the system of structural equations that describe the sales and advertising of the two competing products:

- ❖ $-\text{Log}(\text{Salesf}_t) + 0 \cdot \text{Log}(\text{Salesnf}_t) + \beta_1 \text{Log}(\text{Adf}_t) + \beta_2 \text{Log}(\text{Adnf}_t) + \gamma_1 \text{Log}(\text{PDI}_t) + \gamma_2 \text{Log}(\text{Price}_t) + \epsilon_{1t} = 0$
- ❖ $0 \cdot \text{Log}(\text{Salesf}_t) - \text{Log}(\text{Salesnf}_t) + \beta_3 \text{Log}(\text{Adf}_t) + \beta_4 \text{Log}(\text{Adnf}_t) + \gamma_4 \text{Log}(\text{PDI}_t) + \gamma_5 \text{Log}(\text{Price}_t) + \epsilon_{2t} = 0$
- ❖ $-\text{Log}(\text{Salesf}_t) + \beta_5 \text{Log}(\text{Salesnf}_t) + \beta_6 \text{Log}(\text{Adf}_t) + 0 \cdot \text{Log}(\text{Adnf}_t) + 0 \cdot \text{Log}(\text{PDI}_t) + 0 \cdot \text{Log}(\text{Price}_t) + \epsilon_{3t} = 0$
- ❖ $\beta_7 \text{Log}(\text{Salesf}_t) - \text{Log}(\text{Salesnf}_t) + 0 \cdot \text{Log}(\text{Adf}_t) + \beta_8 \text{Log}(\text{Adnf}_t) + 0 \cdot \text{Log}(\text{PDI}_t) + 0 \cdot \text{Log}(\text{Price}_t) + \epsilon_{4t} = 0$

The reduced form equations for this system are:

- ❖ $\text{Log}(\text{Salesf}_t) = \alpha_1 \text{Log}(\text{PDI}_t) + \alpha_2 \text{Log}(\text{Price}_t) + \epsilon_{1t}$
- ❖ $\text{Log}(\text{Salesnf}_t) = \alpha_4 \text{Log}(\text{PDI}_t) + \alpha_5 \text{Log}(\text{Price}_t) + \epsilon_{2t}$
- ❖ $\text{Log}(\text{Adf}_t) = \alpha_7 \text{Log}(\text{PDI}_t) + \alpha_8 \text{Log}(\text{Price}_t) + \epsilon_{3t}$
- ❖ $\text{Log}(\text{Adnf}_t) = \alpha_{10} \text{Log}(\text{PDI}_t) + \alpha_{11} \text{Log}(\text{Price}_t) + \epsilon_{4t}$

Under certain circumstances the structural parameters may be estimated without testing the model; However if the structural equations are unidentified or over-identified, estimation procedures are debatable.

IDENTIFICATION OF THE MODEL USING SAS:

Commands used in SAS:

```
proc syslin data= sasuser.data1 3sls;
endogenous salesf salesnf adf adnf;
```

```

instruments pdi price;
model salesf=adf adnf pdi price;
model salesnf=adf adnf pdi price;
model adf=salesf salesnf;
model adnf=salesf salesnf;
run;

```

Output Obtained in SAS:

The SYSLIN Procedure
Two-Stage Least Squares Estimation

Model **FIRST**

Dependent Variable salesf

Label Salesf

Analysis of Variance

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	2	0.606875	0.303437	151.93	<.0001
Error	10	0.019972	0.001997		
Corrected Total	12	0.878311			

Root MSE 0.04469 **R-Square** 0.96814

Dependent Mean 3.04292 **Adj R-Sq** 0.96177

Coeff Var 1.46865

Warning: The model is not of full rank. Least Squares solutions for the parameters are not unique. Certain statistics will be misleading. A reported degree of freedom of 0 or B means the estimate is biased.

The following parameters have been set to zero. These variables are a linear combination of other variables as shown.

Intercept = +0.1120 * salesnf +0.008022 * adf +0.1813 * PDI
 adnf = -0.0426 * salesnf -0.2579 * adf -0.1353 * PDI
 Price = -0.4315 * salesnf -0.1227 * adf +0.2284 * PDI

Parameter Estimates

Variable	DF	Parameter Estimate	Standard Error	t Value	Pr > t	Variable Label
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Parameter Estimates

Variable	DF	Parameter Estimate	Standard Error	t Value	Pr > t	Variable Label
Intercept	0	0	.	.	.	Intercept
salesnf	0	-0.20442	0.440859	-0.46	0.6528	salesnf
Adf	0	0.702683	0.245282	2.86	0.0168	Adf
Adnf	0	0	.	.	.	Adnf
PDI	0	1.164988	0.389058	2.99	0.0135	PDI
Price	0	0	.	.	.	Price

The SYSLIN Procedure

Two-Stage Least Squares Estimation

Model **SECOND**

Dependent Variable Salesnf

Label Salesnf

Analysis of Variance

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	2	0.089641	0.044821	29.24	<.0001
Error	10	0.015329	0.001533		
Corrected Total	12	0.100975			

Root MSE 0.03915 **R-Square** 0.85397

Dependent Mean 3.32805 **Adj R-Sq** 0.82476

Coeff Var 1.17643

Warning: The model is not of full rank. Least Squares solutions for the parameters are not unique. Certain statistics will be misleading. A reported degree of freedom of 0 or B means the estimate is biased.

The following parameters have been set to zero. These variables are a linear combination of other variables as shown.

Intercept = -0.0238 * adf + 0.2406 * PDI - 0.2595 * Price
 salesf = +0.7608 * adf + 1.0568 * PDI + 0.4737 * Price
 adnf = -0.2458 * adf - 0.1579 * PDI + 0.0988 * Price

Parameter Estimates

Variable	DF	Parameter Estimate	Standard Error	t Value	Pr > t	Variable Label
Intercept	0	0	.	.	.	Intercept
salesf	0	0	.	.	.	salesf
Adf	0	-0.28426	0.110870	-2.56	0.0282	adf
adnf	0	0	.	.	.	adnf
PDI	0	0.529357	0.136513	3.88	0.0031	PDI
Price	0	-2.31738	0.895048	-2.59	0.0270	Price

The SYSLIN Procedure
Two-Stage Least Squares Estimation

Model **THIRD**

Dependent Variable adf

Label adf

Analysis of Variance

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	2	0.605773	0.302886	82.51	<.0001
Error	10	0.036709	0.003671		
Corrected Total	12	1.030716			

Root MSE 0.06059 **R-Square** 0.94286

Dependent Mean -0.47325 **Adj R-Sq** 0.93144

Coeff Var -12.80253

Parameter Estimates

Variable	DF	Parameter Estimate	Standard Error	t Value	Pr > t	Variable Label
Intercept	1	-9.86833	6.371185	-1.55	0.1524	Intercept
salesnf	1	1.418795	1.419381	1.00	0.3411	salesnf
salesf	1	1.535777	0.545511	2.82	0.0183	Salesf

The SYSLIN Procedure
Two-Stage Least Squares Estimation

Model **FOURTH**
Dependent Variable Adnf
Label Adnf

Analysis of Variance

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	2	0.043281	0.021641	37.60	<.0001
Error	10	0.005755	0.000576		
Corrected Total	12	0.086330			

Root MSE 0.02399 **R-Square** 0.88264
Dependent Mean -0.49102 **Adj R-Sq** 0.85916
Coeff Var -4.88565

Parameter Estimates

Variable	DF	Parameter Estimate	Standard Error	t Value	Pr > t	Variable Label
Intercept	1	1.739523	2.522642	0.69	0.5062	Intercept
salesnf	1	-0.31649	0.561998	-0.56	0.5857	salesnf

IDENTIFICATION OF THE MODEL USING STATA:

The conventional order condition (necessary but not sufficient) pertaining to single-equation estimation with instrumental variables is satisfied by counting included endogenous and excluded exogenous variables in the equation. The sufficient rank condition pertains to the rank of the matrix of instruments.

At present, Stata's **reg3** command does not check to see that the conditions for identification of a structural system are satisfied, and produces estimation results. The **checkreg3** command allows you to verify that these results are meaningful by checking to see that the rank condition is satisfied for each of the N equations in the system. Unless the rank condition is satisfied for each equation in the system,

the system is unidentified. Although unusual for a system to satisfy the order condition without satisfying the rank condition, it can occur.

Regression Results with CheckReg3:

```
. checkreg3 ( logsalesf logadf logadnf logpdi logprice) ( logsalesnf logadf lo
> gadnf logpdi logprice) ( logadf logsalesf logsalesnf) ( logadnf logsalesf logs
> alesnf)
```

Endogenous coefficients matrix

	logsalesf	logsalesnf	logadf	logadnf
logsalesf	-1			
logsalesnf	0	-1		
logadf	.5	.5	-1	
logadnf	.5	.5	0	-1

Exogenous coefficients matrix

	logpdi	logprice
logsalesf	.5	.5
logsalesnf	.5	.5
logadf	0	0
logadnf	0	0

Eq 1 fails rank condition for identification

Eq 2 fails rank condition for identification

Eq 3 fails rank condition for identification

Eq 4 fails rank condition for identification

Rank deficiency: System is not identified

SUMMARY AND CONCLUSION:

From the above studied model on sales and advertising expenditure of cigarettes, it can be seen that each of the equations violate the Rank Condition necessary for identification as some of the parameters in each of the structural equations do not have unique solutions. In the simultaneous model specified above the following structural parameters remain unidentified:

$$\beta_1(adf), \beta_2(adnf), \beta_3(adf), \beta_4(adnf), \gamma_1(pdi), \gamma_2(price), \gamma_4(pdi), \gamma_5(prices)$$

Since, the model is not of full rank, the Least Squares solutions for the parameters are not unique. The estimates obtained will be misleading. There is possible biasedness of estimates.

Despite its limitations, the simultaneous equation regression model can be successfully applied to advertising problems and may aid in shaping managerial decisions. Estimates of the unidentified structural parameters have been developed by two stage least squares regression, but the significance of this estimation has not been determined.

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Appendix

Commands Used:

SAS:

```
proc syslin data= sasuser.data1 3sls;  
endogenous salesf salesnf adf adnf;  
instruments pdi price;  
model salesf=adf adnf pdi price;  
model salesnf=adf adnf pdi price;  
model adf=salesf salesnf;  
model adnf=salesf salesnf;  
run;
```

STATA:

```
Checkreg3 (LogSalesf LogAdf LogAdnf LogPDI LogPrice) (LogSalesnf LogAdf LogAdnf LogPDI LogPrice) (LogAdf  
LogSalesf LogSalesnf) (LogAdnf LogSalesf LogSalesnf)
```

Data Sources:

The data used in the study was available from the year 1953 to 1965.

YEAR	LOG(SALESF)	LOG(SALESNF)	LOG(ADF)	LOG(ADNF)	LOG(PDI)	LOG(PRICE)
1953	2.39851	3.50465	-1.26117	-0.28369	3.41653	-0.60906
1954	2.6006	3.45582	-0.90035	-0.37119	3.41876	-0.60906
1955	2.8389	3.42632	-0.62703	-0.43061	3.44491	-0.60206
1956	2.97883	3.38979	-0.43572	-0.44389	3.46147	-0.60033
1957	3.09065	3.3381	-0.34364	-0.55378	3.46451	-0.59176
1958	3.15067	3.30278	-0.34605	-0.53839	3.46304	-0.5986
1959	3.18361	3.30251	-0.3051	-0.54141	3.47986	-0.57349
1960	3.19626	3.2992	-0.33548	-0.53467	3.48502	-0.57675
1961	3.20779	3.29484	-0.34157	-0.54432	3.49358	-0.57675

1962	3.21945	3.27891	-0.36206	-0.54872	3.50804	-0.5784
1963	3.23843	3.2572	-0.28542	-0.5458	3.51834	-0.56543
1964	3.22329	3.20154	-0.29571	-0.54809	3.54063	-0.55596
1965	3.23099	3.21304	-0.31297	-0.49872	3.56335	-0.5391

The data obtained on the sales of filter cigarettes was an aggregation over the following brands: WINSTON, KENT, MARLBORO, HERBERT TAREYTON, VICEROY, L&M and Parliament. The data on sales of non-filter cigarettes include the following brands: Pall Mall, Camel, Lucky Strike, Chester Field, Old Gold and Philip Moris.