

Probability distributions

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- Economists problem: creating a quantifiable hypothesis
- Econometricians problem: validating the hypothesis
- Focus of the problem: an economic RV and it's probability distribution.
- Some concepts in probability: PDs, PDFs, CDs, CDFs

Moment generating functions of PDs/PDFs

Expectation of RVs

- $E(x)$
- For a discrete RV:

$$E(x) = \sum_{i=1}^n x_i \Pr(x_i)$$

where $\Pr(x)$ is the probability distribution of x .

- For a continuous RV:

$$E(x) = \int_{x=-\infty}^{x=\infty} xf(x)dx$$

where $f(x)$ is the probability density function of x .

- Example: bernoulli RV.
 $x = 0, 1$; $\Pr(0) = p$, $\Pr(1) = 1 - p$

$$E(x) = 0 * p + 1 * (1 - p) = 1 - p$$

Testing concepts: expectation of a discrete variable

RV x , can take the following discrete values, each with equal probability:

x	-1	2	5	7	10	11	12	15	20	30
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What is $E(x)$?

Testing concepts: expectation of a discrete variable

With each value taking equal probability, the PD for x is:

x	-1	2	5	7	10	11	12	15	20	30
$x*\text{Pr}(x)$	-0.1	0.2	0.5	0.7	1.0	1.1	1.2	1.5	2.0	3.0

$$E(x) = \sum x*\text{Pr}(x) = 11.1$$

Example: expectation of continuous variables

- Uniform Continuous RV: $x = [L,U]$; $\Pr(x_i) = p = 1/(U - L)$

$$\begin{aligned} E(x) &= \int_L^U \frac{x}{U-L} d(x) \\ &= \frac{1}{U-L} \int_L^U x d(x) = \frac{1}{U-L} \left(\frac{x^2}{2} \right)_L^U \\ &= \frac{U+L}{2} \end{aligned}$$

Example: expectation of continuous variables

- Normal RV: $x = [-\infty, \infty]$; $\Pr(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2}(x-\mu/\sigma)^2}$

$$E(x) = \int_{-\infty}^{\infty} x f(x) d(x)$$

$$= \frac{e^{1/2\sigma^2}}{\sqrt{2\pi}\sigma} \int_{-\infty}^{\infty} x e^{-\frac{1}{2}(x-\mu)^2} d(x)$$

$$\text{Set } y = x - \mu$$

$$E(x) = C \int_{-\infty}^{\infty} y e^{-\frac{1}{2}y^2} d(y) - \mu C \int_{-\infty}^{\infty} e^{-\frac{1}{2}y^2} d(y)$$

$$= C \int_{-\infty}^{\infty} y e^{-\frac{1}{2}y^2} d(y) - \mu \int_{-\infty}^{\infty} f(x) d(x)$$

Example: expectation of normal distribution

- First integral on the RHS:

$$\begin{aligned}\int y e^{-\frac{1}{2}y^2} d(y) &= \left(-e^{-\frac{y^2}{2}}\right)_{-\infty}^{\infty} \\ &= 0\end{aligned}$$

- Then the expectation becomes:

$$\begin{aligned}E(x) &= 0 + \mu \int_{-\infty}^{\infty} f(x) d(x) \\ E(x) &= \mu\end{aligned}$$

Testing concepts: Expectation of a uniform RV

- Uniform Continuous RV: $x = [0, 10]$. What is $E(x)$?

$$\begin{aligned} E(x) &= \int_0^{10} x f(x) d(x) \\ &= \int_0^{10} x \frac{1}{10} d(x) \\ &= \frac{1}{10} \int_0^{10} x d(x) \\ &= \frac{1}{10} \left(\frac{x^2}{2} \right)_0^{10} \\ &= 5 \end{aligned}$$

Expectations of functions of discrete RVs

- Function $g(x)$ of a random variable x is a random variable. Then, $g(x)$ has a probability distribution based on $\Pr(x)$.
- For any **discrete RV**, x , with a known PD, $\Pr(x)$, the expectation of any function $g()$ of x is calculated as:

$$E(g(x)) = \sum_{\min}^{\max} g(x)Pr(x)$$

- For any **continuous RV** y , with a known PDF, $f(y)$, the expectation of any function $g()$ of y is calculated as:

$$E(g(y)) = \int_{\min}^{\max} g(y)f(y)dy$$

Example: $E(x^2)$ for a binary variable

- Bernoulli RV: $x = 0, 1$, $\Pr(x) = p, (1-p)$
- $g(x) = x^2 = 0, 1$, $\Pr(g(x)) = p, (1-p)$
- $E(g(x)) = E(x^2) = 0 \cdot p + 1 \cdot (1-p) = (1-p)$

Testing concepts: binary variable

RV x is binary with the following probability distribution:

x	$\Pr(x)$	x^2	$x^2 \cdot \Pr(x)$
2	0.3	4	1.2
5	0.7	25	17.5

Questions:

- 1 What is $E(x^2)$? 18.7

Example: $E(x^2)$ for a continuous uniform variable

- Uniform Continuous RV: $x = [L,U]$; $f(x_i) = p = 1/(U - L)$
- $g(x) = x^2$

$$\begin{aligned} E(x^2) &= \int_L^U \frac{x^2}{U-L} d(x) \\ &= \frac{1}{U-L} \int_L^U x^2 d(x) = \frac{1}{U-L} \left(\frac{x^3}{3} \right)_L^U \\ &= \frac{U^3 - L^3}{3 * (U - L)} \end{aligned}$$

- Example: $L=0, U=10$; What is $E(x^2)$?
- $\Pr(x_i) = 1/10, E(x^2) = 1000/30 = 33.3333$

Moment generating functions

- For any distribution, there can be a series of “moments” calculated as follows:

$$\text{(Discrete rv)} \quad E(x^i) = \sum_{\min}^{\max} x^i Pr(x)$$

$$\text{(Continuous rv)} \quad E(x^i) = \int_{-\infty}^{\infty} x^i f(x) d(x)$$

- Each moment describes a feature of the distribution.

The unique moments of a distribution

- The moments are functions of the parameters of the distribution.
- Every distribution has as many **unique** moments as parameters.
The remainder of the moments can be expressed as functions of the parameters.
- For example, the bernoulli distribution had $E(x) = E(x^2) = (1-p)$, the probability of success.
- For example, every moment of the normal distribution can be expressed as a function of the first two moments, the mean (μ) and the variance (σ^2).

Numerical tools to describe a distribution

Statistical measures for a distribution

- Measure of location: Mean, mode, median
- Measure of dispersion: Variance, range, quartiles

Measures of location

- **Mean:** An expected value on a random draw from the dataset.
- **Mode:** The value that occurs with the maximum frequency. Easily interpreted for discrete variables. The mode for the continuous RV datasets is interpreted in terms of the “range/set” of values that are most often observed.
- **Median:** The value of the RV at which 50% of the dataset is observed.

- Find the mean of the data: 5, 1, 6, 2, 4:

$$\bar{x} = \frac{\sum x}{n} = \frac{18}{5} = 3.6$$

Examples: Median

- Find the median of: 1, 7, 3, 1, 4, 5, 3.
- First step is to order the data: 1, 1, 3, 3, 4, 5, 7.
- The median is 3, the midway point, for an odd number of data.
- When the data has an even number of points, the median is calculated as the midpoint between the two choices.
- For a dataset: 9, 5, 7, 3, 1, 8, 4, 6, ordered as 1, 3, 4, 5, 6, 7, 8, 9, the median is 5.5.

Pros and cons of location measures

- Typically, all three measures tend to cluster together - the differences are not very large.
- **However** the mean is most sensitive to the presence of **outliers**.
(For example, a day on which a trader places a buy limit order for 100 million shares of Reliance instead of a thousand shares.)
- The median is less sensitive to the mean. It is not influenced by the value of the observations, just their number.
Thus, it can be a more robust measure of location than the mean.

Measures of dispersion

- Range: The difference between the highest and the lowest value of the RV in the dataset.

Example: In data, 3, 7, 2, 1, 8 the range = $8 - 1 = 7$.

- Variance: The value of the RVs as differences from the average value, squared and summed up. It is denoted by $\sigma(x)^2$.

$$\sigma(x) = \frac{\sum x_i - \bar{x}}{(n - 1)}$$

Example: $\bar{x} = 4.2, \sigma(x)^2 = (-1.2^2 + 2.8^2 + -2.2^2 + -3.2^2 + 3.8^2)/4 = 9.7$.

Calculating the data dispersion using \bar{x}, σ^2

- Question: what is the range of values of the RV between which we can find 95% of the data?
- Answer:
 - 1 Upper range value = $\bar{x} + 1.96 * \sigma$
 - 2 Lower range value = $\bar{x} - 1.96 * \sigma$

- $\bar{x} \pm \sigma = 85\%$ of the dataset
The percentage will be larger for more skewed distributions. The percentage will be closer to 70% for distributions that are more symmetric.
- $\bar{x} \pm 2\sigma = 97\%$ of the dataset
- $\bar{x} \pm 3\sigma = 99\%$ of the dataset

Measures of dispersion: Percentiles, Quartiles

- Percentiles: Denoted as p^{th} percentile. The value of RV, x , such that $p\%$ of the dataset falls below the value x , and $(100 - p)\%$ is above.
- Quartiles: A set of three specific percentiles at the 25^{th} , 50^{th} , 75^{th} percentiles. They are the lower, median and upper quartile values.
The median is the 2^{nd} quartile and the 50^{th} percentile.
- Inter-quartile range (IQR): The distance between the lower and the upper quartile values.

Setting up likelihood framework for estimation

Syntax: Population distributions

- Given a population prob. distribution, $P(x)$, the frequency of x in the population is denoted as $f(x)$.
- Given a sample of size n , the frequency of x in the sample is denoted as $\hat{f}(x)$.
- $f(x)$ is a deterministic function of the PD/PDF.
 $\hat{f}(x)$ is a random variable, which is the function of the sample!
- $f(x)$ is always the same for a given x .
 $\hat{f}(x)$ varies depending upon the sample.

Recap on PDFs vs. CDFs

- Theory about distributions focuses on CDFs.

$$F(x) = P(X \leq x)$$

This is well defined, irrespective of the type of rv.

- From the CDFs we can calculate the joint distribution of X, Y as:

$$F(x, y) = P(X \leq x \text{ and } Y \leq y)$$

- From the CDF we calculate the marginal distribution function as:

$$F(y) = P(Y \leq y) = P(X \leq \infty \text{ and } Y \leq y)$$

- If X, Y are independent, the joint distribution function is the product of the marginals:

$$P(X \leq x \text{ and } Y \leq y) = P(X \leq x)P(Y \leq y)$$

Developing a full statistical model

The Bernoulli model

- Question: What is the population frequency of girls among new born children?
- Economists hypothesis: If Y_i is the gender of child i , then:
 - 1 Y_i are independent across all i .
 - 2 Y_i come from an *identical* distribution
 - 3 Y_i are Bernoulli distributed, with $p = \theta$
 - 4 θ will take values between 0 and 1.

(Note: Are all these reasonable assumptions?)

- Econometrician's task: Find the correct θ
- Get a data: same dataset as the UK dataset of newborn children where $P(Y = \text{boy}) = 0.513\%$ and $N \sim 715,000$ observations.

The likelihood function approach

- We analyse how **probable** the different outcomes observed are for a given θ .
- Start from scratch: if we know θ , then for a sample of size N , can we write the probability of observing $Y_1 = y_1, Y_2 = y_2, \dots, Y_N = y_N$?
- Assuming independence (assumption #1), it is:

$$\begin{aligned}P(Y_1 \leq y_1, \dots, Y_N \leq y_N) &= P(Y_1 \leq y_1) \dots P(Y_N \leq y_N) \\&= \prod_{i=1}^N P(Y_i \leq y_i) \\&= \prod_{i=1}^N \theta^{y_i} (1 - \theta)^{(1-y_i)} \\&= \theta^{\sum_1^N y_i} (1 - \theta)^{\sum_1^N (1-y_i)}\end{aligned}$$

$$\text{But } \bar{y} = \frac{1}{N} \sum_1^N y_i$$

$$P(Y_1 \leq y_1, \dots, Y_N \leq y_N) = \theta^{n\bar{y}} (1 - \theta)^{n(1-\bar{y})}$$

The likelihood function, L

- $P(Y_1 \leq y_1, Y_2 \leq y_2, \dots, Y_N \leq y_N | \theta) = f_\theta(y_1, y_2, \dots, y_N)$
Here, the probability is a function of a known θ . The (y_1, y_2, \dots, y_N) is a set of outcomes in a sample of size N generated by parameter θ .
- We flip this around to asking: given an N -sized sample, can we use the likelihood of a given value of θ to have generated the observed sample of size N ? Or,
- What is the likelihood of θ given the sample:

$$L_{Y_1, Y_2, \dots, Y_N}(\theta) = f_\theta(Y_1, Y_2, \dots, Y_N)$$

Here, Y_1, \dots, Y_N is a fixed set of random observations in the dataset.

- The likelihood function depends **only** on the observations.

The likelihood function for the Bernoulli model

- Given the sample, we say:

$$L_{(Y_1, \dots, Y_N)}(\theta) = \theta^{\bar{Y}}(1 - \theta)^{(1 - \bar{Y})^N}$$

- For the Bernoulli model, $L(\theta)$ depends only on \bar{Y}
- Therefore, \bar{Y} becomes a sufficient statistic for θ .
- Estimation is about how to find the best value of θ given \bar{Y} :
Find θ such that L is maximised.
- Which brings us to optimisation theory given a function.